

## Stable estimation of the probability density function of intensity from photon frequency counts

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A numerical solution to the problem of estimating the probability density of integrated intensity,  $P(W)$ , given a measured histogram of photon counts is described. The solution has a minimum norm in a Hilbert space that is weighted according to a prior estimate of  $P(W)$  and is stable by virtue of a simple but powerful method of regularization. Additional stabilization is achieved by restricting the number of input photon-count frequencies.

### 1. INTRODUCTION

A common problem in the measurement of photon-count statistics is to infer the probability density function (PDF) of intensity from those of the photon counts. Let  $P(W)$ , where  $W \geq 0$ , be the PDF of integrated intensities, and let  $p(n)$ , where  $n = 0, 1, 2, \dots$ , be the PDF of photon counts resulting from light distributed by  $P(W)$ . In the semiclassical approximation, the two PDF's are related through the Poisson transform<sup>1</sup> by the relation

$$p(n) = \int_0^\infty (W^n/n!) e^{-W} P(W) dW. \quad (1.1)$$

A histogram of measured photon-count probabilities is modeled as

$$y(n) = p(n) + \eta, \quad (1.2)$$

where the term  $\eta$  is a zero mean and arbitrary variance-covariance structured random error. Saleh<sup>2</sup> showed that an analytic solution for  $P(W)$  given  $p(n)$  is possible in principle. Solutions based on analytic methods have been proposed<sup>3,4</sup> and have been demonstrated to work when  $p(n)$  is known exactly. When  $p(n)$  is estimated from a measured histogram and substituted into these formulas, even small sampling errors cause wild fluctuations in the resulting estimates for  $P(W)$ . It will be shown that a more stable numerical estimate of  $P(W)$  can be formed given the vector of  $y(n)$ , a prior estimate of the basis PDF  $P_0(W)$ , and its corresponding photon-count PDF  $p_0(n)$ .

### 2. HILBERT-SPACE FORMULATION

In order to consider  $P(W)$  to be a member of a suitably defined Hilbert space, rewrite Eq. (1.1) as

$$p(n) = \int_0^\infty P(W) [P_0(W) W^n/n!] e^{-W} P_0^{-1}(W) dW. \quad (2.1)$$

Here,  $P(W)$  is a member of the Hilbert space  $G = L^2(0, \infty, Q_0^{-1})$ , where  $Q_0(W) = e^W P_0(W)$ . The inner product of the

two functions  $f(W)$  and  $g(W)$  in  $G$  is defined as

$$\langle f, g \rangle_G = \int_0^\infty f(W) g(W) Q_0^{-1}(W) dW, \quad (2.2)$$

so that, with

$$g_n(W) = P_0(W) W^n/n!,$$

Eq. (2.1) is written as

$$p(n) = \langle P, g_n \rangle_G, \quad n = 0, 1, \dots$$

Let the transformation  $A: G \rightarrow H = R^{N+1}$  (the Euclidean  $N+1$  space) be that which associates with any  $P(W)$  in  $G$  the vector  $\mathbf{p} = [p(0), p(1), \dots, p(N)]^T$  ( $T$  denotes the transpose of the vector). The adjoint mapping is  $A^T: H \rightarrow G$  and assigns to each  $\mathbf{h} = [h_0, h_1, \dots, h_N]^T$  in  $R^{N+1}$  the function

$$A^T \mathbf{h} = P_0(W) \sum_{n=0}^N h_n W^n/n!. \quad (2.3)$$

The data vector  $\mathbf{y}$  is in  $H$ , and there will be exact solutions to the equation

$$\mathbf{y} = \mathbf{A} \mathbf{g}. \quad (2.4)$$

In particular, the minimum norm solution  $\mathbf{g}_{mn}$  given by Eq. (A3) is

$$\mathbf{g}_{mn} = \mathbf{A}^T (\mathbf{A} \mathbf{A}^T)^{-1} \mathbf{y}. \quad (2.5)$$

The transformation  $\mathbf{A} \mathbf{A}^T$  associates with each  $\mathbf{h}$  in  $H$  the vector of data values  $\mathbf{y}$ , and, in particular, the  $m$ th value is

$$\begin{aligned} y(m) &= \int_0^\infty \left[ P_0(W) \sum_{n=0}^N h_n W^n/n! \right] e^{-W} (W^m/m!) dW \\ &= \sum_{n=0}^N h_n \left[ \int_0^\infty P_0(W) (W^{m+n}/m!n!) e^{-W} dW \right] \\ &= \sum_{n=0}^N h_n \{ [(m+n)!/m!n!] p_0(m+n) \}. \end{aligned}$$

In matrix notation, this becomes

$$y = P_0 h, \tag{2.6}$$

where the mapping  $AA^T$  is represented by the matrix  $P_0$  in Eq. (2.6). Numerical estimates of  $P(W)$  are obtained by taking

$$h = P_0^{-1} y \tag{2.7}$$

and

$$g_{mn}(W) = P_0(W) \sum_{n=0}^N h_n W^n / n! \tag{2.8}$$

The matrix  $P_0$  is frequently ill conditioned in practice, so that small changes in the data  $y$  produce large changes in the numerical values of  $g_{mn}$ . Stabilization of the estimates requires use of regularization methods.

### 3. REGULARIZED SOLUTION

As is discussed in Appendix A, the regularized solution of Eq. (2.6) compromises between an exact solution and one that controls  $\|g\|$ . This solution, given by Eq. (A4), is

$$g_r = (A^T A + \epsilon I)^{-1} A^T y. \tag{3.1}$$

One cannot obtain numerical values from Eq. (3.1) since  $A^T A$  is an infinite dimensional operator. To obtain a solution in terms of a possible matrix inversion, clear the inverse from Eq. (3.1) and rearrange the terms to get

$$g_r = (1/\epsilon) A^T (y - A g_r) \\ = A^T h_r \tag{3.2}$$

To find  $h_r$ , again multiply Eq. (3.1) by  $(A^T A + \epsilon I)$  and then by  $A$ . Because  $P_0 = AA^T$  is invertible, it follows that

$$h_r = (P_0 + \epsilon I)^{-1} y. \tag{3.3}$$

One obtains numerical estimates by substituting Eq. (3.3) into Eq. (3.2) and, as in Eq. (2.8),

$$g_r(W) = P_0(W) \sum_{n=0}^N h_{rn} W^n / n! \tag{3.4}$$

One sees that the regularization involves only the adding of  $\epsilon > 0$  to the main diagonal of the matrix  $P_0$ . Even a small value of  $\epsilon$  ( $\approx 10^{-4}$ ) is sufficient in some cases to stabilize  $\|g\|$ . So long as  $\epsilon$  is greater than the computing machine's zero and not so large as to dominate the other terms in the matrix ( $P_0 + \epsilon I$ ), a stable solution should be obtained. Generally, the regularized solution  $g_r$  is not consistent with the data [i.e., it is not a true solution of the matrix equation (2.4)]. By using Eqs. (3.2) and (3.3), one sees that, for

$$A g_r = (AA^T)(AA^T + \epsilon I)^{-1} y,$$

only as  $\epsilon \rightarrow 0$  does  $g_r$  approach data consistency.

The regularization presented here is sometimes referred to as Tychonov regularization<sup>5</sup> or as Miller regularization.<sup>6</sup> In fact, the basic idea of involving a measure of the norm of the solution in the function being minimized occurs in a variety of areas. This approach can also be viewed as an extension of the Hoerl-Kennard ridge regression estimator<sup>7</sup> to the case of infinite dimensional regression. The use of the prior basis PDF  $P_0(W)$  in the construction of  $G$  also has application in problems of spectral estimation and signal processing.<sup>8,9</sup>

### 4. NUMERICAL EXAMPLE

We used a computer simulation to determine if a negative exponential PDF can be estimated from its Poisson transform pair, the Bose-Einstein PDF; i.e., for

$$p(n) = \mu^n / (1 + \mu)^{n+1} \tag{4.1}$$

Eq. (1.1) can be solved to obtain

$$P(W) = 1/\mu \exp(-W/\mu). \tag{4.2}$$

A sample of 10,000 counting intervals was simulated in a computer by using Eq. (4.1) and the parameter  $\mu = 4$ . The number of photon counts  $n$  was accumulated in a frequency histogram shown in Table 1. The data fit was accomplished by using a Bose-Einstein prior PDF with  $\mu = 1$ . It follows that the following matrices and functions are:

$y = [p(n) + \eta]$ , i.e., the vector of simulated photon-count probabilities,

$$P_0 = \left[ \frac{(m+n)!}{m!n!} p_0(m+n) \right] = \left[ \frac{(m+n)!}{m!n!} \left(\frac{1}{2}\right)^{m+n} \right],$$

$$P_0(W) = e^{-W}.$$

Equations (3.3) and (3.4) were solved to obtain regularized estimates of the intensity PDF  $g_r(W)$ . Figure 1 shows these

Table 1. Summary of Photon-Count Simulation<sup>a</sup>

Count	Frequency	Count	Frequency	Count	Frequency
0	1984	12	140	24	13
1	1602	13	98	25	6
2	1280	14	100	26	5
3	1012	15	64	27	6
4	832	16	52	28	2
5	647	17	44	29	0
6	501	18	43	30	5
7	452	19	37	31	0
8	338	20	24	32	0
9	236	21	18	33	3
10	223	22	16	34	2
11	199	23	13	-	-

<sup>a</sup> Overflow counts, 3; total sample size, 10,000.

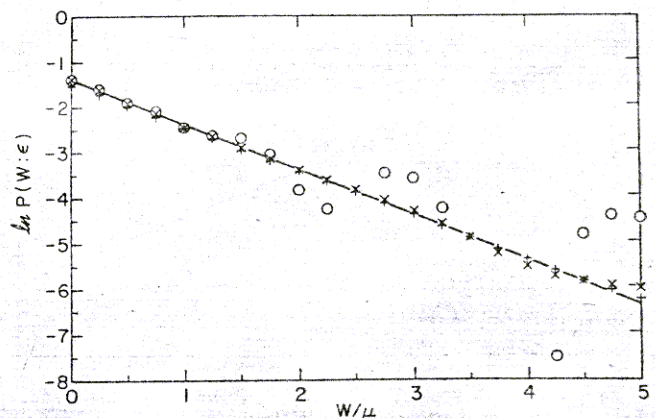


Fig. 1. Regularized estimates of  $P(W)$  from simulated Bose-Einstein photon counts. Simulation parameters: sample size, 10,000;  $\mu = 4$ . Regularization parameters:  $\mu_0 = 1$ ,  $\epsilon$ : O, 0.0001; X, 0.01; +, 0.1. The first 29 photon-count frequencies were used to compute  $P_0$ .

estimates for differing values of  $\epsilon$  on a logarithmic scale for the intensity PDF and on a normalized scale for  $W$ . Estimates are defined as admissible if they are positive and less than one. For  $\epsilon = 0$ , there are no admissible estimates; however, one can see that an admissible estimation is obtained for all values of  $W$  by using  $\epsilon = 10^{-4}$ . The exponentially increasing values of  $\epsilon$  generally improve the estimation until too large a value biases the estimate of  $P(W)$  at the origin and at its tails.

## 5. DISCUSSION

Other approaches to the estimation of  $P(W)$ , such as those of Simar<sup>10</sup> and of Tucker,<sup>11</sup> focus on the asymptotic properties of the probability density estimators. For finite data these estimators are discrete probabilities supported on a finite number of points (about  $N/2$ ). These estimators are good (as Simar shows) for cases in which the compounding PDF is essentially a sum of point masses, i.e., a finite Poisson mixture. For the problem considered in this paper, such estimators are inappropriate. One advantage of Eq. (3.4) is that it is a continuous function of intensity.

Finite sampling errors in estimating the photon-count frequencies by  $y(n)$  are also an important consideration in regularizing the  $P_0$  matrix. Matrix condition can be improved by restricting the input photon counts only to reliable estimates. Reducing the dimension of  $P_0$ , and hence also the basis vector  $h$ , is not severe for smooth functions, such as the exponential given in Section 4. This truncation, however, may reduce the range of  $W/\mu$  over which accurate estimates of intensity may be obtained.

The importance of matrix conditioning of  $P_0$  is illustrated in Figs. 2 and 3. When the correct prior parameter is used, the resulting estimates are accurate only through a small range of intensities for all values of  $\epsilon$ , as shown by Fig. 2. Further conditioning of  $P_0$  is still needed to obtain a satisfactory estimation. This is accomplished in Fig. 3, in which only relative photon-count frequencies greater than 0.1 are used. This suggests consideration of a measure for the condition of  $P_0$ , namely, the ratio of the smallest to the largest eigenvalue. The conditioning measure would also aid in making the choice for a minimum acceptable value for  $\epsilon$ .

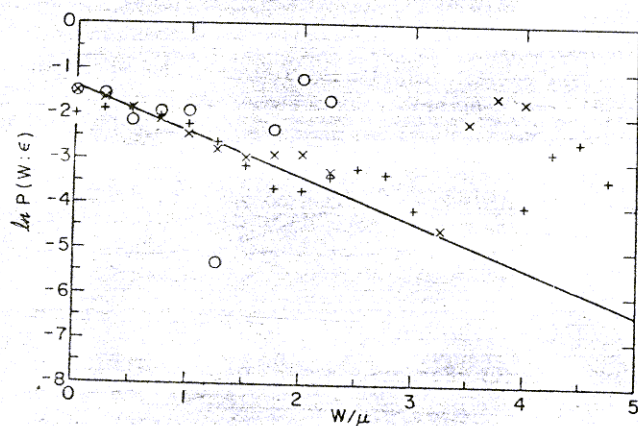


Fig. 2. Regularized estimates of  $P(W)$  from simulated Bose-Einstein photon counts. Simulation parameters: sample size, 10,000;  $\mu = 4$ . Regularization parameters:  $\mu_0 = 4$ ;  $\epsilon$ : O, 0.0001; X, 0.01; +, 0.1. The first 29 photon-count frequencies were used to compute  $P_0$ .

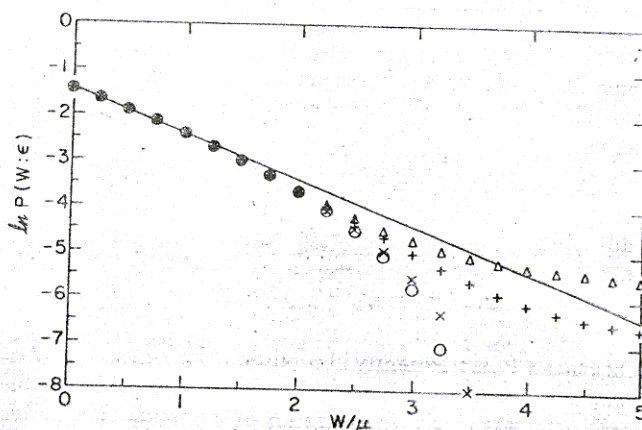


Fig. 3. Regularized estimates of  $P(W)$  from simulated Bose-Einstein photon counts. Simulation parameters: sample size, 10,000;  $\mu = 4$ . Regularization parameters:  $\mu_0 = 4$ ;  $\epsilon$ : O, 0.0; X, 0.0001; +, 0.0005;  $\Delta$ , 0.001. The first four photon-count frequencies were used to compute  $P_0$ .

No positivity constraints have been applied, nor has the integral of the estimated PDF been forced to equal one. This leaves open the possibility of using Eq. (3.4) iteratively with a succession of prior estimates  $P_0(W)$ , each obtained by clipping the previous regularized estimate  $P_r(W)$  to make it positive and calling that a new prior estimate. Note that this procedure should iteratively discover a better basis for which the resulting regularized estimate is more positive.

## APPENDIX A

Let  $G$  and  $H$  be arbitrary (possibly infinite dimensional) Hilbert spaces and the operator  $A: G \rightarrow H$  be any continuous linear transformation. For the vector  $h_0$ , a fixed but arbitrary member of  $H$ , the equation

$$Ag = h_0 \quad (A1)$$

may have no solutions, one unique solution, or infinitely many solutions. In general, one considers the minimum-norm least-squares solution, defined as the unique  $g$  in  $G$  of minimum norm among all  $g$  for all  $\|Ag - h_0\|$ , as minimum. Denote this solution by  $g_{mnl}$ . Two special cases are important.

*Case 1. A is one to one.* There is a unique  $g$  minimizing Eq. (A1) so that the minimum-norm least-squares solution is the least-squares solution; i.e.,

$$g_{mnl} = g_{ls} = (A^T A)^{-1} A^T h_0, \quad (A2)$$

where  $A^T$  is the adjoint transformation mapping  $H$  into  $G$ .

*Case 2. A is onto H.* The transformation  $A^T$  is one to one, and there is at least one solution to Eq. (A1). The minimum-norm least-squares solution is now the minimum-norm solution, or

$$g_{mnl} = g_{mn} = A^T (A A^T)^{-1} h_0. \quad (A3)$$

If small changes in  $h_0$  produce large changes in the minimum-norm solution  $g_{mn}$ , then one can regularize by replacing the minimum-norm solution with a least-squares solution to a slightly different problem, which is formulated below.

Assume that  $A$  is onto  $H$  so that Eq. (A1) has exact solutions. Let  $K$  be the Hilbert space of ordered pairs  $\mathbf{k} = (\mathbf{g}, \mathbf{h})$ , that is,  $K = G \oplus H$ . The inner product is

$$\begin{aligned} \langle \mathbf{k}_1, \mathbf{k}_2 \rangle &= \langle (\mathbf{g}_1, \mathbf{h}_1), (\mathbf{g}_2, \mathbf{h}_2) \rangle \\ &= \alpha \langle \mathbf{g}_1, \mathbf{g}_2 \rangle + (1 - \alpha) \langle \mathbf{h}_1, \mathbf{h}_2 \rangle, \end{aligned}$$

so

$$\|\mathbf{k}\|^2 = \|(\mathbf{g}, \mathbf{h})\|^2 = \alpha \|\mathbf{g}\|^2 + (1 - \alpha) \|\mathbf{h}\|^2.$$

Define the transformation  $B: G \rightarrow K$  by the equation

$$B\mathbf{g} = (\mathbf{g}, A\mathbf{g}).$$

Then the adjoint transformation  $B^T: K \rightarrow G$  satisfies the relation

$$B^T(\mathbf{g}, \mathbf{h}) = \alpha\mathbf{g} + (I - \alpha)A^T\mathbf{h}.$$

Given a fixed  $\mathbf{h}_0$  in  $H$ , let  $\mathbf{k}_0 = (\mathbf{0}, \mathbf{h}_0)$ . Unless  $\mathbf{h}_0 = \mathbf{0}$ ,  $\mathbf{k}_0$  is not in the range of  $B$ . Then find a least-squares solution of the equation

$$B\mathbf{g} = \mathbf{k}_0.$$

From the above, it is

$$\begin{aligned} \mathbf{g}_{ls} &= (B^T B)^{-1} B^T \mathbf{k}_0 \\ &= [(1 - \alpha)A^T A + \alpha I]^{-1} (1 - \alpha)A^T \mathbf{h}_0 \\ &= (A^T A + \epsilon I)^{-1} A^T \mathbf{h}_0, \quad \epsilon = \alpha/(1 - \alpha). \end{aligned} \quad (A4)$$

The  $\mathbf{g}_{ls}$  minimizes  $\|B\mathbf{g} - \mathbf{k}_0\|$ , so  $\mathbf{g}_{ls}$  also minimizes the quantity

$$(1 - \alpha) \|A\mathbf{g} - \mathbf{h}_0\|^2 + \alpha \|\mathbf{g}\|^2.$$

The least-squares solution of Eq. (A4) is called the regularized solution in that a strict solution to Eq. (A1) is relaxed in order to gain control over the quantity  $\|\mathbf{g}\|$ .

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